Lecture 7: Hypothesis testing and classification

CME/STATS 195

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October 18, 2018

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Hypothesis testing

Hypothesis testing can answer questions:

- Is the measured quantity equal to/higher/lower than a given threshold? e.g. is the number of faulty items in an order statistically higher than the one guaranteed by a manufacturer?
- Is there a difference between two groups or observations? e.g. Do treated patient have a higher survival rate than the untreated ones?
- Is the level of one quantity related to the value of the other quantity? e.g. Is hyperactivity related to eating sugar? Is lung cancer related to smoking?

To perform a hypothesis test you need to:

- 1. Define the null and alternative hypotheses.
- 2. Choose level of significance α .
- 3. Pick and compute test statistics.
- 4. Compute the p-value.
- 5. Check whether to reject the null hypothesis by comparing p-value to α .
- 6. Draw conclusion from the test.

Null and alternative hypotheses

The null hypothesis (H_0): A statement assumed to be true unless it can be shown to be incorrect beyond a reasonable doubt. This is something one usually attempts to disprove or discredit.

The alternate hypothesis (H_1): A claim that is contradictory to H0 and what we conclude when we reject H0.

HO and H1 are on purporse set up to be contradictory, so that one can collect and examine data to decide if there is enough evidence to reject the null hypothesis or not.

	Contrising Won	ds in Str	2757105	my brain hurt
failed	to reject th	e null	hypothecis	
1			rypotnesis	Alex



Student's t-test

- William Gosset (1908), a chemist at the Guiness brewery.
- Published in Biometrika under a pseudonym Student.
- Used to select best yielding varieties of barley.
- Now one of the standard/traditional methods for hypothesis testing.

Among the typical applications:

- Comparing population mean to a constant value
- Comparing the means of two populations
- Comparing the slope of a regression line to a constant

In general, used when the test statistic would follow a normal distribution if the value of a scaling term in the test statistic were known.

Distribution of the t-statistic



If $X_i \sim \mathcal{N}(\mu, \sigma^2)$, the empirical estimates for mean and variance are: $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ and $s^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

The t-statistic is:

$$T = \frac{\bar{X} - mu}{s/\sqrt{n}} \sim t_{\nu=n-1}$$

p-value

- p-value is the probability of obtaining the same or "more extreme" event than the one observed, assuming the null hypothesis holds (is true).
- A small p-value, typically < 0.05, indicates strong evidence against the null hypothesis; in this case you can reject the null hypothesis.
- A large p-value, > 0.05, indicates weak evidence against the null hypothesis; in this case, you do NOT reject the null hypothesis.





p-values should NOT be used a "ranking"/"scoring" system for your hypotheses

Two-sided test of the mean

Is the mean flight arrival delay statistically equal to 0?

Test the null hypothesis:

 $H_0: \mu = \mu_0 = 0$ $H_a: \mu \neq \mu_0 = 0$ where μ is where μ is the average arrival delay.



```
library(tidyverse)
library(nycflights13)
mean(flights$arr_delay, na.rm = T)
```

[1] 6.895377

Is this statistically significant?

```
( tt = t.test(x=flights$arr_delay, mu=0, alternative="two.sided" ) )
```

```
##
## One Sample t-test
##
## data: flights$arr_delay
## t = 88.39, df = 327340, p-value < 2.2e-16
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## 6.742478 7.048276
## sample estimates:
## mean of x
## 6.895377</pre>
```



The function t.test returns an object containing the following components:

names(tt) "parameter" ## [1] "statistic" "p.value" "conf.int" "estimate" ## [6] "null.value" "alternative" "method" "data.name" # The p-value: tt\$p.value *##* [1] 0 *#* The 95% confidence interval for the mean: tt\$conf.int ## [1] 6.742478 7.048276 ## attr(,"conf.level")

[1] 0.95

	_

One-sided test of the mean

One-sided can be more powerful, but the intepretation is more difficult.

Test the null hypothesis:

$$H_0: \mu = \mu_0 = 0$$

 $H_a: \mu < \mu_0 = 0$



Is the average delay 5 or is it lower?

```
( tt = t.test(x=flights$arr_delay, mu=5, alternative="less" ) )
```

```
##
## One Sample t-test
##
## data: flights$arr_delay
## t = 24.296, df = 327340, p-value = 1
## alternative hypothesis: true mean is less than 5
## 95 percent confidence interval:
## -Inf 7.023694
## sample estimates:
## mean of x
## 6.895377
```

Failure to reject is not acceptance of the null hypothesis.



Testing difference between groups

Is the average arrival delay the same for the winter and summer?

Test the null hypothesis:

$$H_0: \mu_a = \mu_b$$
$$H_a: \mu_a \neq \mu_b$$

where μ_a mean arr delay in the winter and μ_b is the mean arr delay in the summer.

t.test(x, y)

Seasonal differences in flight delay

```
flights %>%
   mutate(season = cut(month, breaks = c(0,3,6,9,12))) %>%
ggplot(aes(x = season, y = arr_delay)) + geom_boxplot (alpha=0.1) +
    xlab("Season") + ylab("Arrival delay")
```

Warning: Removed 9430 rows containing non-finite values (stat_boxplot).



Seasonal differences in flight delay

```
flights %>%
  filter(arr_delay < 120) %>%
  mutate(season = cut(month, breaks = c(0, 3, 6, 9, 12))) %>%
  ggplot(aes(x = season, y = arr_delay)) + geom_boxplot (alpha=0.01) +
xlab("Season") + ylab("Arrival delay")
```



Testing seasonal differences in flight delay

```
flights.winter = filter(flights, month %in% c(1,2,3))
flights.summer = filter(flights, month %in% c(7,8,9))
t.test(x=flights.winter$arr_delay, y=flights.summer$arr_delay)
```

```
##
## Welch Two Sample t-test
##
## data: flights.winter$arr_delay and flights.summer$arr_delay
## t = -2.4383, df = 161250, p-value = 0.01476
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.9780344 -0.1063691
## sample estimates:
## mean of x mean of y
## 5.857851 6.400052
```

Exercise

- Go to the "Lec7_Exercises.Rmd" file, which can be downloaded from the class website under the Lecture tab.
- Complete Exercise 1.

Classification

- **Classification** is a supervised methood which deals with prediction outcomes or response variables that are qualitative, or categorical.
- The task is to classify or assign each observation to a category or a class.
- Examples of classification problems include:
 - predicting what medical condition or disease a patient has base on their symptoms,
 - determining cell types based on their gene expression profiles (single cell RNA-seq data).
 - detecting fraudulent transactions based on the transaction history

Logistic Regression

Logistic Regression

- Logistic regression is actually used for **classification**, and not regression tasks, $Y \in \{0, 1\}.$
- The name **regression** comes from the fact that the method **fits a linear** function to a continuous quantity, the log odds of the response.

$$p = P[Y = 1 \mid X]$$
$$\log\left(\frac{p}{1-p}\right) = X\beta = \beta_0 + \beta_1^T x$$

• The method performs **binary classification** (k = 2), but can be generalized to handle k > 2 classes (multinomial logistic regression).

$$g(p) = \log\left(\frac{p}{1-p}\right), \qquad (\text{ logit a link function})$$
$$g^{-1}(\eta) = \frac{1}{1+e^{-\eta}}, \qquad (\text{ logistic function })$$
$$\eta = X\beta, \qquad (\text{ linear predictor })$$

$$E[Y] = P[Y = 1 | X = x] \quad (\text{ probability of}$$
$$= p = g^{-1}(\eta)$$
$$= \frac{1}{1 + e^{-X\beta}}$$

)

outcome)





$\begin{array}{ccc} & \beta_0 = 0, \ \beta_1 = -2 \\ \hline & \beta_0 = 0, \ \beta_1 = -1 \\ \hline & \beta_0 = 0, \ \beta_1 = 1 \\ \hline & \beta_0 = 0, \ \beta_1 = 2 \end{array}$

Grad School Admissions

Suppose we would like to predict students' admission to graduate school based on their GRE, GPA, and the rank of their undergraduate institution.

```
admissions <- read_csv("https://stats.idre.ucla.edu/stat/data/binary.csv")</pre>
```

```
## Parsed with column specification:
## cols(
##
    admit = col_integer(),
    gre = col_integer(),
##
##
    gpa = col_double(),
##
    rank = col_integer()
## )
```

admissions

##	# A t	ibble	: 400	9 x 4	
##	ac	dmit	gre	gpa	rank
##	<]	int> <:	int>	<dbl></dbl>	<int></int>
##	1	Θ	380	3.61	3
##	2	1	660	3.67	3
##	3	1	800	4	1
##	4	1	640	3.19	4
##	5	Θ	520	2.93	4
##	6	1	760	3	2
##	7	1	560	2.98	1
##	8	Θ	400	3.08	2
##	9	1	540	3.39	3
##	10	Θ	700	3.92	2
##	#	. with	390	more	rows



summary(admissions)

##	admit	gre	gpa	rank	
##	Min. :0.0000	Min. :220.0	Min. :2.260	Min. :1.000	
##	1st Qu.:0.0000	1st Qu.:520.0	1st Qu.:3.130	1st Qu.:2.000	
##	Median :0.0000	Median :580.0	Median :3.395	Median :2.000	
##	Mean :0.3175	Mean :587.7	Mean :3.390	Mean :2.485	
##	3rd Qu.:1.0000	3rd Qu.:660.0	3rd Qu.:3.670	3rd Qu.:3.000	
##	Max. :1.0000	Max. :800.0	Max. :4.000	Max. :4.000	
sap	ply (admissions, s	sd)			

##	admit	gre	gpa	rank
##	0.4660867	115.5165364	0.3805668	0.9444602

Check that there are observations included in each subgroup, and whether the data is balanced:

with(admissions, table(admit, rank))

rank ## ## admit 1 2 3 4 ## 0 28 97 93 55 ## 1 33 54 28 12

Logistic Regression in R

- In R logistic regression can be done using a function glm().
- glm stands for Generalized Linear Model.
- The function can fit many other regression models. Use **?glm** to learn more.
- For cases with k > 2 classes, multinom() function from nnet package can be used. To see how go over this example.

Note that currently the column 'admit' and 'rank' in admissions are integers.

```
sapply(admissions, class)
```

admit rank gre gpa ## "integer" "integer" "numeric" "integer"

We convert the two columns to factors.

```
admissions <- mutate(admissions,
 admit = factor(admit, levels = c(0, 1), labels = c("rejected", "admitted")),
 rank = factor(rank, levels = 1:4)
admissions
```

```
## # A tibble: 400 x 4
     admit
##
                gre
                     gpa rank
##
     <fct>
              <int> <dbl> <fct>
                380 3.61 3
## 1 rejected
## 2 admitted
                660 3.67 3
## 3 admitted
                800 4
                         1
## 4 admitted
                640 3.19 4
                520 2.93 4
## 5 rejected
## 6 admitted
                760 3
                         2
## 7 admitted
                560 2.98 1
## 8 rejected
               400 3.08 2
## 9 admitted
                540 3.39 3
## 10 rejected
               700 3.92 2
## # ... with 390 more rows
```

Split data

Divide data into train and test set so that we can evaluate the model accuracy later on. Here we use 60%-20%-20% split.

```
set.seed(78356)
n <- nrow(admissions)</pre>
idx <- sample(1:n, size = n)</pre>
train.idx <- idx[seq(1, floor(0.6*n))]</pre>
valid.idx <- idx [seq(floor(0.6*n)+1, floor(0.8*n))]
train <- admissions[train.idx, ]</pre>
valid <- admissions[valid.idx, ]</pre>
test <- admissions[-c(train.idx, valid.idx), ]</pre>
nrow(train)
```

[1] 240

```
nrow(valid)
## [1] 80
nrow(test)
## [1] 80
```



Fitting a logistic regression model

```
logit_fit <- glm(</pre>
    admit ~ gre + gpa + rank, data = train, family = "binomial")
```

• The first argument,

formula = admit \sim gre + gpa + rank, specifies the linear predictor part, $\eta = X\beta$.

• You need to set the family to family = "binomial" equivalent to choosing a logistic regression, i.e. using a logit link function $g(\cdot)$ in a GLM model.

Logistic regression coefficients for continuous predictors (covariates) give the log fold change in the odds of the outcome corresponding to a unit increase in the predictor.

$$\beta_{cont} = \log \left(\frac{P[Y = 1 \mid X_{cont} = x + 1]}{P[Y = 1 \mid X_{cont} = x]} \right)$$

Categorical features (factors) are first converted to indicator variables and then the model fits separate coefficients for each level of the factor. Coefficients corresponding to a specific indicator variable give the increase/decrease in the log odds of the outcome in case the observation is recorded with that level.

$$\beta_{factor} = \log\left(\frac{P[Y=1 \mid X_{fac} = L]}{P[Y=1 \mid X_{fac} \neq L]}\right)$$

```
coef(logit_fit)
```

rank2 rank3 ## (Intercept) gre gpa ## -2.662567353 0.000921435 0.658045298 -0.510004503 -1.560051191 rank4 ## ## -1.129252168

- For every unit increase in gre, the log odds of admitted (versus rejected) increases by \approx 9.214349810^{-4}.
- For every unit increase in gpa, the log odds increases by ≈ 0.6580453 .
- There are three coefficients for the rank variable, e.g. a student attending a college with rank 2, one with rank 1 (base level), has the log admission odds decreased by \approx -0.5100045.

You can get the confidence intervals for the coefficients with the confint() fuinction

```
confint(logit_fit)
```

Waiting for profiling to be done...

##		2.5 %	97.5 %
##	(Intercept)	-5.595691918	0.172732111
##	gre	-0.001778273	0.003647635
##	gpa	-0.181398218	1.522814525
##	rank2	-1.289858306	0.260377700
##	rank3	-2.483360377	-0.677844965
##	rank4	-2.140151201	-0.167386365

The 95% CI are away from zero which indicates significance.

summary(logit_fit)

```
##
## Call:
\#\# glm(formula = admit ~ gre + gpa + rank, family = "binomial",
##
      data = train)
##
## Deviance Residuals:
##
      Min
                1Q Median
                                  3Q
                                         Max
## -1.4795 -0.9377 -0.7004 1.1883
                                      2.0539
##
## Coefficients:
##
                Estimate Std. Error z value Pr(>|z|)
## (Intercept) -2.6625674 1.4651841 -1.817 0.069183 .
## gre
               0.0009214 0.0013789
                                   0.668 0.503979
## gpa
             0.6580453
                         0.4329230
                                   1.520 0.128510
## rank2 -0.5100045 0.3935431 -1.296 0.194999
## rank3 -1.5600512 0.4583036 -3.404 0.000664
                                                    * * *
## rank4
              -1.1292522 0.5002488 -2.257 0.023984 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 309.52 on 239 degrees of freedom
## Residual deviance: 289.83 on 234 degrees of freedom
## AIC: 301.83
##
## Number of Fisher Scoring iterations: 4
```



Rank variable effect is given with three different coefficients.

We can sse wald.test() function from the aod package to test the overall effect of 'rank'.

```
# install.packages(aod)
library(aod)
wald.test(b = coef(logit_fit), Sigma = vcov(logit_fit), Terms = 4:6)
## Wald test:
## ------
##
```

```
## Chi-squared test:
## X2 = 14.0, df = 3, P(> X2) = 0.0029
```

- b supplies the coefficients,
- Sigma supplies the variance covariance matrix of the error terms,
- Terms indices of the coefficients to be tested; here 4, 5, and 6, corresponding to 'rank'.

The p-value indicates that the overall effect of rank is statistically significant.

Fitted values

library(modelr)
head(train %>% add_predictions(logit_fit, var = "log_odds"))

##	#	A tibble:	6 X 5	5		
##		admit	gre	gpa	rank	log_odds
##		<fct></fct>	<int></int>	<dbl></dbl>	<fct></fct>	<dbl></dbl>
##	1	rejected	640	3.67	3	-1.22
##	2	admitted	700	3.52	4	-0.830
##	3	rejected	400	3.35	3	-1.65
##	4	rejected	580	3.51	2	-0.328
##	5	admitted	640	3.19	4	-1.10
##	6	admitted	580	3.58	1	0.228

##	# /	A tibble:	240 x	8					
##		admit	gre	gpa	rank	admit_odds	admit_prob	admit_pred	admit_pre
##	-	<fct></fct>	<int></int>	<dbl></dbl>	<fct></fct>	<dbl></dbl>	<dbl></dbl>	<fct></fct>	<fct></fct>
##	1	rejected	640	3.67	3	-1.22	0.228	rejected	rejected
##	2	admitted	700	3.52	4	-0.830	0.304	rejected	rejected
##	3	rejected	400	3.35	3	-1.65	0.161	rejected	rejected
##	4	rejected	580	3.51	2	-0.328	0.419	rejected	rejected
##	5	admitted	640	3.19	4	-1.10	0.249	rejected	rejected
##	6	admitted	580	3.58	1	0.228	0.557	admitted	admitted
44	- 7	raiaatad	EGO	2 26	2	1 50	0 100	raiaatad	raiaatad

ed2

HH	1	rejecteu	500	3.30 3	-1.50	0.103	rejecteu	rejecteu
##	8	rejected	460	3.77 3	-1.32	0.211	rejected	rejected
##	9	admitted	560	2.98 1	-0.186	0.454	rejected	rejected
##	10	rejected	580	3.02 2	-0.651	0.343	rejected	rejected
##	#.	with 230	9 more	e rows				



Predictions

Predictions can be computed using predict() function, with the argument type = "response". Otherwise, the default will compute predictions on the scale of the linear predictors.

```
# Must have the same column names as the variables in the model
new_students <- data.frame(</pre>
    gre = c(670, 790, 550),
    gpa = c(3.56, 4.00, 3.87),
    rank = factor(c(1, 2, 2)))
# The output is the probability of admissions for each of the new students.
new students <- new students %>%
 mutate(
    admit_odds = predict(logit_fit, newdata = new_students),
    admit_pred = factor(admit_odds < 0, levels = c(TRUE, FALSE),</pre>
                   labels = c("rejected", "admitted"))
new students
```

##		gre	gpa	rank	admit_odds	admit_pred
##	1	670	3.56	1	0.2974353	admitted
##	2	790	4.00	2	0.1875430	admitted
##	3	550	3.87	2	-0.1191473	rejected

Multiple models

##	# A	A tibbl	e: 80	x 8						
##		admit	gre	gpa	rank	admit_odd	ds_fit1	admit_odds	s_fit2	admit_fit1
##		<fct></fct>	<int></int>	<dbl></dbl>	<fct></fct>		<dbl></dbl>		<dbl></dbl>	<fct></fct>
##	1	reje…	340	2.92	3		-1.99		-1.41	rejected
##	2	reje…	660	3.31	4		-1.01		-1.03	rejected
##	3	admi	300	2.84	2		-1.03		-0.389	rejected
##	4	reje…	500	4	3		-1.13		-1.41	rejected
##	5	reje…	780	3.87	4		-0.526		-1.03	rejected
##	6	reje…	600	3.63	3		-1.28		-1.41	rejected
##	7	reje…	540	3.78	4		-0.807		-1.03	rejected
##	8	admi…	800	3.74	1		0.536		0.163	admitted
##	9	admi…	800	3.43	2		-0.178		-0.389	rejected
##	10	admi…	740	2.97	2		-0.536		-0.389	rejected
##	#	wit	:h 70 n	nore ro	ows, a	nd 1 more	variab]	<pre>Le: admit_t</pre>	fit2 <1	fct>



Evaluating accuracy

Confusion Matrix for model 1
(confusion_matrix_fit1 <- table(true = valid\$admit, pred = valid\$admit_fit1</pre>

pred
true rejected admitted
rejected 56 3
admitted 16 5

Confusion Matrix for model 2
(confusion_matrix_fit2 <- table(true = valid\$admit, pred = valid\$admit_fit2</pre>

##	p	ored	
##	true	rejected	admitted
##	rejected	57	2
##	admitted	16	5

Accuracy for model 1
(accuracy_fit1 <- sum(diag(confusion_matrix_fit1))/sum(confusion_matrix_fit))</pre>

[1] 0.7625

Accuracy for model 2
(accuracy_fit2 <- sum(diag(confusion_matrix_fit2))/sum(confusion_matrix_fit))</pre>

[1] 0.775

We choose a simpler model logit_fit2

))		
))		
1))		
2))		

Expected logit fit2 performance

Performance of our chosen model, logit fit2 can be evaluated on the testset

```
test <- test %>%
    mutate(
        admit_odds = predict(logit_fit2, newdata = test),
        admit_pred = factor(admit_odds < 0,</pre>
                              levels = c(TRUE, FALSE),
                              labels = c("rejected", "admitted"))
(test confusion matrix <- table(pred = test$admit, true = test$admit pred))
##
             true
              rejected admitted
## pred
    rejected
                     48
##
                               9
##
     admitted
                     15
                               8
(test_accuracy <- sum(diag(test_confusion_matrix))/sum(test_confusion_matrix))</pre>
## [1] 0.7
```

So, you should expect your model accuracy to be around 0.7 for a new dataset you collect later.

Exercise

- Go to the "Lec7_Exercises.Rmd" file, which can be downloaded from the class website under the Lecture tab.
- Complete Exercise 2.

Random Forest

Random Forest

- Random Forest is an ensemble learning method based on classification and **regression trees, CART**, proposed by Breinman in 2001.
- RF can be used to perform **both classification and regression**.
- RF models are robust as they combine predictions calculated from a large number of decision trees (a forest).
- Details on RF can be found in Chapter 8 of ISL and Chapter 15 ESL; also a good write-up can also be found here

Decision trees

- Cool visualization explaining what decision trees are: link
- Example of decision trees



Tree bagging Algorithm

Suppse we have an input data matrix, $X \in \mathbb{R}^{N \times p}$ and a response vector, $Y \in \mathbb{R}^N$. For b = 1, 2, ..., B:

1. Generate a random subset of the data (X_b, Y_b) contatining n < Nobservations sampled with replacement.

2. Train a decision tree T_b on (X_b, Y_b)

3. Predict the outcome for N - n unseen (complement) samples (X'_h, Y'_h)

Afterwards, combine predictions from all decision trees and compute the average predicted outcome.

Averaging over a collection of decision trees makes the predictions more stable.

Decision trees for bootrap samples







b = 2





0





Source: Chapter 8 ESL

Random Forest Characteristics

- Random forests differ in only one way from tree bagging: it uses a modified tree learning algorithm sometimes called **feature bagging**.
- At each candidate split in the learning process, only a random subset of the features is included in a pool from which the variables can be selected for splitting the branch.
- Introducing randomness into the candidate splitting variables, reduces correlation between the generated trees.





Random Forest: Ensemble of Trees

Source: link

Wine Quality

UCI ML Repo includes two datasets on red and white variants of the Portuguese "Vinho Verde" wine. The datasets contain information on physicochemical and sensory characteristics of the wine quality score.

We will use the white wines dataset to classify wines according to their quality classes.

```
url <- 'https://archive.ics.uci.edu/ml/machine-learning-databases/wine-quality/winequality-white</pre>
wines <- read.csv(url, sep = ";")</pre>
head(wines, 6)
```

##		fixed.acidity volat	ile.acidity	citric.aci	id residu	ual.su	ugar chlor	ides
##	1	7.0	0.27	0.3	36		20.7 0	.045
##	2	6.3	0.30	0.3	34		1.6 0	.049
##	3	8.1	0.28	0.4	40		6.9 0	.050
##	4	7.2	0.23	0.3	32		8.5 0	.058
##	5	7.2	0.23	0.3	32		8.5 0	.058
##	6	8.1	0.28	0.4	40		6.9 0	.050
##		free.sulfur.dioxide	total.sulf	ur.dioxide	density	рН	sulphates	alcohc
##	1	45		170	1.0010	3.00	0.45	8.
##	2	14		132	0.9940	3.30	0.49	9.
##	3	30)	97	0.9951	3.26	0.44	10.
##	4	47	,	186	0.9956	3.19	0.40	9.
##	5	47	,	186	0.9956	3.19	0.40	9.
##	6	30		97	0.9951	3.26	0.44	10.
##		quality						
##	1	6						
##	2	6						
##	3	6						
##	4	6						
##	5	6						
##	6	6						

Class Frequency

```
table(wines$quality)
```

3 4 5 6 7 8 9 ## 20 163 1457 2198 880 175 5

```
ggplot(wines, aes(x = quality)) +
  geom_bar() + theme_classic() +
  ggtitle("Barplot for Quality Scores")
```





The classes are ordered and not balanced (munch more normal wines than excellent/poor ones).

6

quality

4

8

To make things easier, we will wines into "good", "average" and "bad" categories.

The new classes will be more balanced, and it will be easier to fit the model.

```
qualClass <- function(quality) {
    if(quality > 6) return("good")
    if(quality < 6) return("bad")
    return("average")
}
wines <- wines %>%
    mutate(taste = sapply(quality, qualClass),
        taste = factor(taste, levels = c("bad", "average", "good")))
head(wines)
```

##		fixed.ad	cidity	volat	ile.aci	dity	citric	c.ac:	id	residu	Jal	. sı	ugar	chlor	ides
##	1		7.0			0.27		0.3	36				20.7	Θ	.045
##	2		6.3			0.30		0.3	34				1.6	Θ	.049
##	3		8.1			0.28		0.4	40				6.9	Θ	.050
##	4		7.2			0.23		0.3	32				8.5	Θ	.058
##	5		7.2			0.23		0.3	32				8.5	Θ	.058
##	6		8.1			0.28		0.4	40				6.9	Θ	.050
##		free.su	lfur.di	ioxide	total.	sulfu	ur.dio>	kide	de	nsity		рН	sulp	hates	alcoho
##	1			45				170	1	.0010	З.	00		0.45	8
##	2			14				132	0	.9940	З.	30		0.49	9
##	3			30				97	0	.9951	З.	26		0.44	10
##	4			47				186	0	.9956	З.	19		0.40	9
##	5			47				186	0	.9956	З.	19		0.40	9
##	6			30				97	0	.9951	3.	26		0.44	10
##		quality	tast	ce											
##	1	6	averaç	ge											
##	2	6	averaç	ge											
##	3	6	averaç	ge											
##	4	6	averaç	ge											
##	5	6	averaç	ge											
##	6	6	averaç	ge											

ad" categories. e model.

ol .8 .5 .1 .9 .9 table(wines\$quality)



```
ggplot(wines, aes(x = taste)) +
  geom_bar() + theme_classic() +
  ggtitle("Barplot for Quality Scores")
```







Splitting data

We include 60% of the data in a train set and the remaining into a test set.

```
set.seed(98475)
idx <- sample(nrow(wines), 0.6 * nrow(wines))</pre>
train <- wines[idx, ]</pre>
test <- wines[-idx, ]</pre>
dim(train)
## [1] 2938
                13
dim(test)
## [1] 1960
                13
```

Random Forest in R

In R there is a convenient function randomForest from randomForest package.

```
# install.packages("randomForest")
library(randomForest)
wines_fit_rf <- randomForest(</pre>
    taste \sim . - quality, data = train,
    mtry = 5, ntree = 500, importance = TRUE)
```

- Note that in the formula 'taste ~ . quality' means we include all features EXCEPT for 'quality' (the response variable).
- mtry the number of variables randomly sampled as candidates at each split. Defaults: for classification – \sqrt{p} and for regression – p/3, where p is number of all variables in the model.
- ntree the number of trees in the forest.
- importance whether importance of predictors be computed.

Observe, that RF is good at distinguishing "bad" wines from "good" wines, but still struggles when it comes to "average" wines.

```
wines fit rf
```

```
##
## Call:
## randomForest(formula = taste ~ . - quality, data = train, mtry = 5, ntree = 500, import
##
                 Type of random forest: classification
                       Number of trees: 500
##
## No. of variables tried at each split: 5
##
##
          OOB estimate of error rate: 31.31%
## Confusion matrix:
##
           bad average good class.error
## bad
           681
                  272
                       15
                             0.2964876
                  966 135
## average 219
                             0.2681818
## good
           20
                  259 371
                             0.4292308
```

Model Accuracy

- You should always evaluate your model's performance on a test set, which was set aside and not observed by the method at all.
- In case of RF, performance on train and test set should be similar; this is because the method averages predictions computed by individual trees for observations unseen by the tree.
- Inspect the confusion matrix to asses the model accuracy.

```
(confusion_matrix <- table(</pre>
    true = test$taste, pred = predict(wines_fit_rf, newdata = test)))
```

##		pred		
##	true	bad	average	good
##	bad	482	181	9
##	average	149	669	60
##	good	13	143	254

(accuracy_rf <- sum(diag(confusion_matrix)) / sum(confusion_matrix))</pre>

[1] 0.7168367

https://stats.stackexchange.com/questions/197827/how-to-interpret-mean-decrease-in-accuracy-and-mean-decrease-gini-in-random-fore

Look at variable importance: importance(wines_fit_rf)

##		bad	average	good	MeanDecreaseAccuracy
##	fixed.acidity	30.15194	30.17027	29.82500	51.71162
##	volatile.acidity	64.10513	51.51792	57.95579	90.28951
##	citric.acid	28.54081	32.93660	31.90320	46.52323
##	residual.sugar	29.23441	35.39843	27.38350	56.88708
##	chlorides	36.06739	26.80210	39.22203	49.98833
##	free.sulfur.dioxide	37.74602	35.26059	29.29246	57.27752
##	total.sulfur.dioxide	25.84618	23.53196	34.53854	45.42788
##	density	26.92925	28.25958	29.45976	43.55052
##	рН	33.72925	31.09405	42.54602	56.16315
##	sulphates	29.16720	28.56807	30.09379	47.44873
##	alcohol	81.11168	36.20917	66.60965	94.30226
##		MeanDecre	easeGini		
##	fixed.acidity	-	133.9582		
##	volatile.acidity		205.1542		
##	citric.acid	-	L43.4607		
##	residual.sugar	-	159.3942		
##	chlorides	-	158.9609		
##	free.sulfur.dioxide	-	173.0973		
##	total.sulfur.dioxide	-	160.1464		
##	density	-	186.5196		
##	рН	-	162.8367		
##	sulphates	-	138.5101		
##	alcohol		258.7888		

What seems to be the conclusion? What are the characteristics that are predictive of the wine quality score?



Exercise

- Go to the "Lec7_Exercises.Rmd" file, which can be downloaded from the class website under the Lecture tab.
- Complete Exercise 3.